

# Large matchings in uniform hypergraphs and the conjectures of Erdős and Samuels

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## Abstract

In this paper we study conditions which guarantee the existence of perfect matchings and perfect fractional matchings in uniform hypergraphs. We reduce this problem to an old conjecture by Erdős on estimating the maximum number of edges in a hypergraph when the (fractional) matching number is given, which we are able to solve in some special cases using probabilistic techniques. Based on these results, we obtain some general theorems on the minimum  $d$ -degree ensuring the existence of perfect (fractional) matchings. In particular, we asymptotically determine the minimum vertex degree which guarantees a perfect matching in 4-uniform and 5-uniform hypergraphs. We also discuss an application to a problem of finding an optimal data allocation in a distributed storage system.

## 1 Introduction

A  $k$ -uniform hypergraph or a  $k$ -graph for short, is a pair  $H = (V, E)$ , where  $V := V(H)$  is a finite set of vertices and  $E := E(H) \subseteq \binom{V}{k}$  is a family of  $k$ -element subsets of  $V$  called edges. Whenever convenient we will identify  $H$  with  $E(H)$ . A *matching* in  $H$  is a set of disjoint edges of  $H$ . The number of edges in a matching is called *the size* of the matching. The size of the largest matching in a  $k$ -graph  $H$  is denoted by  $\nu(H)$ . A matching is *perfect* if its size equals  $|V|/k$ .

A *fractional matching* in a  $k$ -graph  $H = (V, E)$  is a function  $w : E \rightarrow [0, 1]$  such that for each  $v \in V$  we have  $\sum_{e \ni v} w(e) \leq 1$ . Then  $\sum_{e \in E} w(e)$  is the size of  $w$ . The size of the largest fractional

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matching in a  $k$ -graph  $H$  is denoted by  $\nu^*(H)$ . If  $\nu^*(H) = n/k$ , or equivalently, for all  $v \in V$  we have  $\sum_{e \ni v} w(e) = 1$ , then we call  $w$  *perfect*.

The determination of  $\nu^*(H)$  is a linear programming problem. Its dual problem is to find a minimum *fractional vertex cover*  $\tau^*(H) = \sum_{v \in V} w(v)$  over all functions  $w : V \rightarrow [0, 1]$  such that for each  $e \in E$  we have  $\sum_{v \in e} w(v) \geq 1$ . Let  $\tau(H)$  be the minimum number of vertices in a vertex cover of  $H$ . Then, for every  $k$ -graph  $H$ , by the Duality Theorem,

$$\nu(H) \leq \nu^*(H) = \tau^*(H) \leq \tau(H). \quad (1)$$

Given a  $k$ -graph  $H$  and a set  $S \in \binom{V}{d}$ ,  $0 \leq d \leq k - 1$ , we denote by  $\deg_H(S)$  the number of edges in  $H$  which contain  $S$ . Let  $\delta_d := \delta_d(H)$  be the minimum  $d$ -degree of  $H$ , which is the minimum  $\deg_H(S)$  over all  $S \in \binom{V}{d}$ . Note that  $\delta_0(H) = |E(H)|$ . In this paper we study the relation between the minimum  $d$ -degree  $\delta_d(H)$  and the matching numbers  $\nu(H)$  and  $\nu^*(H)$ .

**Definition 1.1** Let integers  $d, k, s$ , and  $n$  satisfy  $0 \leq d \leq k - 1$ , and  $0 \leq s \leq n/k$ . We denote by  $m_d^s(k, n)$  the minimum  $m$  so that for an  $n$ -vertex  $k$ -graph  $H$ ,  $\delta_d(H) \geq m$  implies that  $\nu(H) \geq s$ . Equivalently,

$$m_d^s(k, n) - 1 = \max\{\delta_d(H) : |V(H)| = n \text{ and } \nu(H) \leq s - 1\}.$$

Furthermore, letting  $s$  be a *real* number satisfying  $0 \leq s \leq n/k$ , define  $f_d^s(k, n)$  as the minimum  $m$  so that  $\delta_d(H) \geq m$  implies that  $\nu^*(H) \geq s$ . Equivalently,

$$f_d^s(k, n) - 1 = \max\{\delta_d(H) : |V(H)| = n \text{ and } \nu^*(H) < s\}.$$

Observe that trivially, for  $\lceil s \rceil \leq n/k$ ,

$$f_d^s(k, n) \leq m_d^{\lceil s \rceil}(k, n). \quad (2)$$

We are mostly interested in the case  $s = n/k$  (i.e. when matchings are perfect) in which we suppress the superscript in the notation  $m_d^{n/k}(k, n)$  and  $f_d^{n/k}(k, n)$ . Thus, writing  $m_d(k, n)$ , we implicitly require that  $n$  is divisible by  $k$ .

Problems of this type have a long history going back to Dirac [4] who in 1952 proved that minimum degree  $n/2$  implies the existence of a Hamiltonian cycle in graphs. Therefore, for  $d \geq 1$ , we refer to the extremal parameters  $m_d(k, n)$  and  $f_d(k, n)$  as to *Dirac-type thresholds*. When  $k = 2$ , an easy argument shows that  $m_1(2, n) = n/2$ . For  $k \geq 3$ , an exact formula for  $m_{k-1}(k, n)$  was obtained in [25]. For a fixed  $k \geq 3$  and  $n \rightarrow \infty$  it yields the asymptotics  $m_{k-1}(k, n) = \frac{n}{2} + O(1)$ . As far as perfect fractional matchings are concerned, it was proved in [23] that  $f_{k-1}(k, n) = \lceil n/k \rceil$  for  $k \geq 2$ , which is a lot less than  $m_{k-1}(k, n)$  when  $k \geq 3$ . For more results on Dirac-type thresholds for matchings and Hamilton cycles see [22].

In this paper, we focus on the asymptotic behavior of  $m_d(k, n)$  and  $f_d(k, n)$  for general, but fixed  $k$  and  $d$ , when  $n \rightarrow \infty$ . For a lower bound on  $m_d(k, n)$  consider first a  $k$ -graph  $H_0 = H_0(k, n)$

(constructed in [25]) with vertex set split almost evenly, that is,  $V(H_0) = A \cup B$ ,  $||A| - |B|| \leq 2$ , and with the edge set consisting of all  $k$ -element subsets of  $V(H_0)$  intersecting  $A$  in an odd number of vertices. We choose the size of  $A$  so that  $|A|$  and  $\frac{n}{k}$  have different parity. Clearly, there is no perfect matching in  $H_0$  and for every  $0 \leq d \leq k - 1$  we have  $\delta_d(H_0) \sim \frac{1}{2} \binom{n-d}{k-d}$ .

Another lower bound on  $m_d(k, n)$  is given by the following well known construction. For integers  $n, k$ , and  $s$ , let  $H_1(s)$  be a  $k$ -graph on  $n$  vertices consisting of all  $k$ -element subsets intersecting a given set of size  $s - 1$ , that is  $H_1(s) = K_n^{(k)} - K_{n-s+1}^{(k)}$ . Observe that  $\nu(H_1(s)) = s - 1$ , while

$$\delta_d(H_1(n/k)) = \binom{n-d}{k-d} - \binom{n-d-n/k+1}{k-d} \sim \left\{ 1 - \left( \frac{k-1}{k} \right)^{k-d} \right\} \binom{n-d}{k-d}.$$

Assume that  $n$  is divisible by  $k$ . Putting  $s = \frac{n}{k}$  and using the  $k$ -graphs  $H_0$  and  $H_1(n/k)$ , we obtain a lower bound

$$m_d(k, n) \geq \max \{ \delta_d(H_0), \delta_d(H_1(\frac{n}{k})) \} + 1 \sim \max \left\{ \frac{1}{2}, 1 - \left( \frac{k-1}{k} \right)^{k-d} \right\} \binom{n-d}{k-d}. \quad (3)$$

On the other hand,  $H_1(\lceil n/k \rceil)$  alone yields a lower bound also on  $f_d(k, n)$ . Indeed, for a real  $s > 0$  we have

$$\nu^*(H_1(\lceil s \rceil)) = \tau^*(H_1(\lceil s \rceil)) \leq \tau(H_1(\lceil s \rceil)) = \lceil s \rceil - 1 < s,$$

and so

$$f_d(k, n) \geq \delta_d(H_1(\lceil \frac{n}{k} \rceil)) + 1 \sim \left\{ 1 - \left( \frac{k-1}{k} \right)^{k-d} \right\} \binom{n-d}{k-d}. \quad (4)$$

It is easy to check that for  $d \geq k/2$  the maximum in the coefficient in (3) equals  $\frac{1}{2}$ . Pikhurko [21] proved, complementing the case  $d = k - 1$ , that indeed we have  $m_d(k, n) \sim \frac{1}{2} \binom{n-d}{k-d}$  also for  $k/2 \leq d \leq k - 2$ ,  $k \geq 4$ .

For  $d < k/2$  the problem seems to be harder and we discuss below the cases  $d \geq 1$  and  $d = 0$  separately. The first result for the range  $1 \leq d < k/2$ ,  $k \geq 3$ , was obtained already in 1981 by Daykin and Häggkvist in [3] who proved that  $m_1(k, n) \leq \left( \frac{k-1}{k} + o(1) \right) \binom{n-1}{k-1}$ . This was generalized to  $m_d(k, n) \leq \left( \frac{k-d}{k} + o(1) \right) \binom{n-d}{k-d}$  for all  $1 \leq d < k/2$  in [10], and, using the ideas from [10], slightly improved in [19] to  $m_d(k, n) \leq \left\{ \frac{k-d}{k} - \frac{1}{k^{k-d}} + o(1) \right\} \binom{n-d}{k-d}$ . For  $k = 4, d = 1$  the latter coefficient is  $\frac{47}{64}$ . In [19], the constant was further lowered to  $\frac{42}{64}$ , still away from the lower bound of  $\frac{37}{64}$ .

It has been conjectured in [15] and again in [10] that the lower bound (3) is achieved at least asymptotically.

**Conjecture 1.1** *For all  $1 \leq d \leq k - 1$ ,*

$$m_d(k, n) \sim \max \left\{ \frac{1}{2}, 1 - \left( \frac{k-1}{k} \right)^{k-d} \right\} \binom{n-d}{k-d}.$$

Hàn, Person, and Schacht in [10] proved Conjecture 1.1 in the case  $d = 1$ ,  $k = 3$  by showing that  $m_1(3, n)$  is asymptotically equal to  $\frac{5}{9} \binom{n-1}{2}$ . Kühn, Osthus, and Treglown [16] and, independently, Khan [13], proved the exact result  $m_1(3, n) = \delta_1(H_1(n/3)) + 1$ . Recently Khan [14] announced that he verified the exact result  $m_1(4, n) = \delta_1(H_1(n/4)) + 1$ , while the asymptotic version,  $m_1(4, n) \sim \frac{37}{64} \binom{n-1}{3}$  follows also from a more general result by Lo and Markström [18].

These exact results, together with (2) and (4), yield that  $f_1(3, n) = m_1(3, n)$  and  $f_1(4, n) = m_1(4, n)$ . Remembering that, on the other hand,  $f_{k-1}(k, n)$  is much smaller than  $m_{k-1}(k, n)$ , one can raise the question about a general relation between  $m_d(k, n)$  and its fractional counterpart  $f_d(k, n)$ . In this paper we answer this question by showing that  $m_d(k, n)$  and  $f_d(k, n)$  are asymptotically equal whenever  $f_d(k, n) \sim c^* \binom{n-d}{k-d}$  for some constant  $c^* > \frac{1}{2}$ , and otherwise  $m_d(k, n) \sim \frac{1}{2} \binom{n-d}{k-d}$ .

**Theorem 1.1** *For every  $1 \leq d \leq k - 1$  if there exists  $c^* > 0$  such that  $f_d(k, n) \sim c^* \binom{n-d}{k-d}$  then*

$$m_d(k, n) \sim \max \left\{ c^*, \frac{1}{2} \right\} \binom{n-d}{k-d}. \quad (5)$$

This result reduces the task of asymptotically calculating  $m_d(k, n)$  to a presumably simpler task of calculating  $f_d(k, n)$ . It seems that, similarly to the integral case, the lower bound in (4) determines asymptotically the actual value of the parameter  $f_d(k, n)$ .

**Conjecture 1.2** *For all  $1 \leq d \leq k - 1$ ,*

$$f_d(k, n) \sim \left\{ 1 - \left( \frac{k-1}{k} \right)^{k-d} \right\} \binom{n-d}{k-d}.$$

Our next result confirms Conjecture 1.2 asymptotically for all  $k$  and  $d$  such that  $1 \leq k - d \leq 4$ . Note that the above mentioned result from [23] shows that Conjecture 1.2 is true for  $d = k - 1$  exactly, that is,  $f_{k-1}(k, n) = \delta_{k-1}(H_1(\lceil \frac{n}{k} \rceil)) + 1$ . We include this case into the statement of Theorem 1.2 for completeness.

**Theorem 1.2** *For every  $k \geq 3$  and  $k - 4 \leq d \leq k - 1$ , we have*

$$f_d(k, n) \sim \left\{ 1 - \left( \frac{k-1}{k} \right)^{k-d} \right\} \binom{n-d}{k-d}.$$

Theorems 1.2 and 1.1 together imply immediately the validity of Conjecture 1.1 in a couple of new instances (as discussed earlier, the first of them has been recently also proved in [14] and [18]).

**Corollary 1.1** *We have*

$$\begin{aligned} m_1(4, n) &\sim \frac{37}{64} \binom{n-1}{3}, & m_2(5, n) &\sim \frac{1}{2} \binom{n-2}{3}, & m_1(5, n) &\sim \frac{369}{625} \binom{n-1}{4} \\ m_2(6, n) &\sim \frac{671}{1296} \binom{n-2}{4}, & m_3(7, n) &\sim \frac{1}{2} \binom{n-3}{4}. \end{aligned}$$

We prove Theorem 1.2 utilizing the following connection between the parameters  $f_d^s(k, n)$  and  $f_0^s(k-d, n-d)$ .

**Proposition 1.1** *For all  $k \geq 3$ ,  $1 \leq d \leq k-1$ , and  $n \geq k$ ,*

$$f_d(k, n) \leq f_0^{n/k}(k-d, n-d).$$

In view of Proposition 1.1, in order to prove Theorem 1.2 we need to estimate  $f_0^s(k-d, n-d)$  with  $s = \frac{n}{k}$ . This is trivial for  $d = k-1$  and so, from now on, we will be assuming that  $d \leq k-2$ . The integral version of this problem has almost as long history as the Dirac-type problem ( $d \geq 1$ ).

Erdős and Gallai [6] determined  $m_0^s(k, n)$  for graphs ( $k = 2$ ). In 1965, Erdős [5] conjectured the following hypergraph generalization of their result.

**Conjecture 1.3** *For all  $k \geq 2$  and  $1 \leq s \leq \frac{n}{k}$ :*

$$m_0^s(k, n) = \max \left\{ \binom{ks-1}{k}, \binom{n}{k} - \binom{n-s+1}{k} \right\} + 1.$$

The lower bound comes from considering again the extremal  $k$ -graph  $H_1(s)$  along with the  $k$ -uniform clique  $K_{ks-1}^{(k)}$  (complemented by  $n-ks+1$  isolated vertices) which, clearly, has no matching of size  $s$ . For more on Erdős' conjecture we refer the reader to the survey paper [7] and a recent paper [9], where the conjecture is proved for  $k = 3$  and  $n \geq 4s$ . In its full generality, the conjecture is still wide open.

We now formulate the fractional version of Erdős' Conjecture. For future references, we switch from  $k$  and  $n$  to  $l$  and  $m$ . Again, the lower bound is yielded by  $H_1(\lceil s \rceil)$  and the complete  $l$ -graph on  $\lceil ls \rceil - 1$  vertices,  $K_{\lceil ls \rceil - 1}^{(l)}$ .

**Conjecture 1.4** *For all integers  $l \geq 2$  and a real  $s$  such that  $0 \leq s \leq m/l$ , we have*

$$f_0^s(l, m) = \max \left\{ \binom{\lceil ls \rceil - 1}{l}, \binom{m}{l} - \binom{m - \lceil s \rceil + 1}{l} \right\} + 1.$$

As a consequence of the Erdős-Gallai theorem from [6], Conjecture 1.4 is asymptotically true for  $l = 2$ . In the next section we establish a result which confirms Conjecture 1.4 asymptotically in the two smallest new instances, but limited to the range  $0 \leq s \leq \frac{m}{l+1}$ . In this range the case  $l = 3$  follows also from the above mentioned result in [9]. It is easy to check that for  $s \leq \frac{m}{l+1} + O(1)$ , the maximum in Conjecture 1.4 is achieved by the second term.

**Theorem 1.3** *For  $l \in \{3, 4\}$ , for all  $d \geq 1$ , and  $s = \frac{m+d}{l+d}$*

$$f_0^s(l, m) \sim \left\{ 1 - \left( 1 - \frac{1}{l+d} \right)^l \right\} \binom{m}{l}$$

where the asymptotics holds for  $m \rightarrow \infty$  with  $d$  fixed.

Theorem 1.3 together with Proposition 1.1 implies Theorem 1.2, which, in turn, together with Theorem 1.1 yields Corollary 1.1. To prove Conjecture 1.1 in full generality, one would need to prove Theorem 1.3 for all  $l$ .

The rest of this paper is organized as follows. In the next section, we prove Theorem 1.3 using as a main tool a probabilistic inequality of Samuels. A proof of Proposition 1.1, and consequently of Theorem 1.2, appears in Section 3. Section 4 contains a proof of Theorem 1.1. Finally, in Section 5, we discuss an application of the fractional version of the Erdős problem in distributed storage allocation. The last section contains concluding remarks and open problems.

## 2 Fractional matchings and probability of small deviations

In this section we prove Theorem 1.3 using a probabilistic approach from [1] based on a special case of an old probabilistic conjecture of Samuels [26]. In fact, we prove a little bit more – see Corollary 2.1 and Remark 2.1 below.

For  $l$  reals  $\mu_1, \dots, \mu_l$  satisfying  $0 \leq \mu_1 \leq \mu_2 \leq \dots \leq \mu_l$  and  $\sum_{i=1}^l \mu_i < 1$ , let

$$P(\mu_1, \mu_2, \dots, \mu_l) = \inf \mathbb{P}(X_1 + \dots + X_l < 1),$$

where the infimum is taken over all possible collections of  $l$  independent nonnegative random variables  $X_1, \dots, X_l$ , with expectations  $\mu_1, \dots, \mu_l$ , respectively. Define

$$Q_t(\mu_1, \dots, \mu_l) = \prod_{i=t+1}^l \left( 1 - \frac{\mu_i}{1 - \sum_{j=1}^t \mu_j} \right)$$

for each  $0 \leq t < l$ .

Note that  $Q_t(\mu_1, \dots, \mu_l)$  is exactly  $\mathbb{P}(X_1 + \dots + X_l < 1)$  when  $X_i$  is identically  $\mu_i$  for all  $i \leq t$ , while  $X_i$  attains the values 0 and  $1 - \sum_{i \leq t} \mu_i$  (with its expectation being  $\mu_i$ ) for all  $i \geq t + 1$ .

The following conjecture was raised by Samuels in [26].

**Conjecture 2.1 ([26])** *For all admissible values of  $\mu_1, \dots, \mu_l$ ,*

$$P(\mu_1, \mu_2, \dots, \mu_l) = \min_{t=0, \dots, l-1} Q_t(\mu_1, \mu_2, \dots, \mu_l).$$

Note that for  $l = 1$  this is Markov's inequality. Samuels proved his conjecture for  $l \leq 4$  in [26, 27].

**Lemma 2.1 ([26, 27])** *The assertion of Conjecture 2.1 holds for all  $l \leq 4$ .*

We next show that for  $\mu_1 = \mu_2 = \dots = \mu_l = x$ , where  $0 < x \leq \frac{1}{l+1}$ , the minimum in Conjecture 2.1 is attained by  $Q_0(\mu_1, \dots, \mu_l)$ .

**Proposition 2.1** For every integer  $l \geq 2$  and every real number  $x$  satisfying  $0 < x \leq \frac{1}{l+1}$ , if  $\mu_1 = \mu_2 = \dots = \mu_l = x$  then

$$\min_{t=0, \dots, l-1} Q_t(\mu_1, \mu_2, \dots, \mu_l) = Q_0(\mu_1, \mu_2, \dots, \mu_l) = (1-x)^l.$$

**Proof:** By definition

$$Q_t(\mu_1, \mu_2, \dots, \mu_l) = \left(1 - \frac{x}{1-tx}\right)^{l-t} = \left(\frac{1-(t+1)x}{1-tx}\right)^{l-t}.$$

We thus have to prove that for  $0 < x \leq \frac{1}{l+1}$  and  $1 \leq t \leq l-1$ ,

$$(1-x)^l \leq \left(\frac{1-(t+1)x}{1-tx}\right)^{l-t}$$

or equivalently that

$$\left(\frac{1}{1-x}\right)^l \geq \left(\frac{1-tx}{1-(t+1)x}\right)^{l-t}.$$

By the geometric-arithmetic means inequality applied to a set of  $l$  numbers,  $t$  of which are equal to 1 and the remaining  $l-t$  equal to the quantity  $\frac{1-tx}{1-(t+1)x}$ , we conclude that

$$\left(\frac{1-tx}{1-(t+1)x}\right)^{l-t} \cdot 1^t \leq \left[\frac{1}{l} \cdot \left(\frac{(1-tx)(l-t)}{1-(t+1)x} + t\right)\right]^l.$$

Thus, it suffices to show that

$$\frac{(1-tx)(l-t)}{1-(t+1)x} + t \leq \frac{l}{1-x}.$$

This is equivalent to

$$(1-x)[(1-tx)(l-t) + t - t(t+1)x] \leq l[1 - (t+1)x],$$

which is equivalent to

$$(1-x)[l - t(l+1)x] \leq l - l(t+1)x,$$

or

$$l - t(l+1)x - lx + t(l+1)x^2 \leq l - l(t+1)x.$$

After dividing by  $x$ , we see that this is equivalent to  $x \leq \frac{1}{l+1}$ , which holds by assumption, completing the proof. ■

Note that when  $s = xm$  and  $x \leq \frac{1}{l+1}$ , the maximum in Conjecture 1.4 is achieved by the second term. We now prove the following, in most part hypothetical, result, which shows how to deduce Conjecture 1.4 in this range from Conjecture 2.1.

**Theorem 2.1** For any  $l \geq 3$  and  $0 < x \leq \frac{1}{l+1}$ , if Conjecture 2.1 holds for  $l$  and  $\mu_1 = \mu_2 = \dots = \mu_l = x$  then

$$f_0^{xm}(l, m) \sim \left\{1 - (1-x)^l\right\} \binom{m}{l}.$$

Combining Theorem 2.1 with Lemma 2.1, we obtain the following corollary which implies Theorem 1.3. (For  $d = 1$ , observe that  $f_0^s(l, m) \sim f_0^s(l, m + 1)$ .)

**Corollary 2.1** *For  $l = 3$ ,  $x \leq 1/4$  and for  $l = 4$ ,  $x \leq 1/5$ , the maximum number of edges in an  $l$ -uniform hypergraph  $H$  on  $m$  vertices with fractional matching number less than  $xm$  is*

$$f_0^{xm}(l, m) \sim \left\{1 - (1 - x)^l\right\} \binom{m}{l}.$$

**Proof of Theorem 2.1:** Let  $H$  be an  $l$ -uniform hypergraph on a vertex set  $V$ ,  $|V| = m$ , and suppose that  $\nu^*(H) < xm$ . By duality, we also have  $\tau^*(H) < xm$ , and hence there exists a weight function  $w : V \rightarrow [0, 1]$  such that  $\sum_{v \in V} w(v) < xm$  and, for every edge  $e$  of  $H$ ,  $\sum_{v \in e} w(v) \geq 1$ . By increasing the weights  $w(v)$  if needed, we may assume that

$$\sum_{v \in V} w(v) = xm.$$

Let  $v_1, \dots, v_l$  be a sequence of random vertices of  $H$ , chosen independently and uniformly at random from  $V$ . For each  $i = 1, \dots, l$  we define a random variable  $X_i = w(v_i)$ . Note that  $X_1, X_2, \dots, X_l$  are independent, identically distributed random variables, where every  $X_i$  attains each of the  $m$  values  $w(v)$  with probability  $1/m$ . (The values of  $w$  for different vertices can be equal, but this is of no importance for us.)

By definition, the expectation  $\mu_i$  of each  $X_i$  is

$$\mu_i = \sum_{v \in V} \frac{1}{m} \cdot w(v) = \frac{xm}{m} = x.$$

Now we can estimate the number of edges of  $H$  as follows. Since for each edge of  $H$  we have  $\sum_{v \in e} w(v) \geq 1$ , the number  $N$  of all  $l$ -element subsets  $S$  of  $V$  with  $\sum_{v \in S} w(v) < 1$  is a lower bound on the number of non-edges of  $H$ . Let  $N_1$  and  $N_2$  be the numbers of all  $l$ -element sequences of vertices of  $V$  and all  $l$ -element sequences of *distinct* vertices of  $V$ , respectively, with the sums of weights strictly smaller than 1. Note that  $N_1 - N_2$  is at most the number of  $l$ -element sequences in which at least one vertex appears twice, thus it is bounded by  $\binom{l}{2} m^{l-1} = O(m^{l-1})$ . As the number of all  $l$ -element subsets of  $V$  is  $\binom{m}{l} = (1 + o(1))m^l/l!$  and  $N = N_2/l!$ , we have

$$\mathbb{P}\left(\sum_{i=1}^l w(v_i) < 1\right) = \frac{N_1}{m^l} \leq \frac{N_2 + O(m^{l-1})}{\binom{m}{l}l!} = (1 + o(1)) \frac{N}{\binom{m}{l}}.$$

If Conjecture 2.1 holds for a given  $l$  then, by Lemma 2.1 and Proposition 2.1,

$$\mathbb{P}\left(\sum_{i=1}^l w(v_i) < 1\right) = \mathbb{P}\left(\sum_{i=1}^l X_i < 1\right) \geq (1 - x)^l,$$

and, consequently,

$$N \geq (1 + o(1))(1 - x)^l \binom{m}{l}.$$

It follows that the number of edges of  $H$  is at most  $(1 + o(1))\{1 - (1 - x)^l\} \binom{m}{l}$ , as needed.  $\blacksquare$



**Remark 2.1** Note that the above proof works as long as the conclusion of Proposition 2.1 holds. One can check using Mathematica that Proposition 2.1 holds for  $l = 3$  and all  $0 < x \leq 0.277$ , as well as for  $l = 4$  and all  $0 < x \leq 0.217$ . Therefore, Corollary 2.1 extends to these broader ranges of  $x$ . For bigger values of  $x$ , e.g., for  $x = 0.3$  when  $l = 3$ , this is not the case anymore, and the above method does not suffice to determine the asymptotic behavior of  $f_0^{xm}(l, m)$ . In fact, using Samuels conjecture in the higher range of  $x$ , one gets a bound on  $f_0^{xm}(l, m)$  which is larger than that in Conjecture 1.4. However, in view of Proposition 1.1, for our main application the case  $x \leq \frac{1}{l+1}$  is just what we need.

### 3 Thresholds for perfect fractional matchings

In this section we present a proof of Proposition 1.1 and then deduce quickly Theorem 1.2.

**Proof of Proposition 1.1:** The outline of the proof goes as follows. We will assume that there is no fractional perfect matching in a  $k$ -graph  $H$  on  $n$  vertices and then show that the neighborhood graph  $H(L)$  in  $H$  of a particular set  $L$  of size  $d$  satisfies  $\nu^*(H(L)) < n/k$ . This will imply that  $\delta_d(H) \leq |H(L)| < f_0^{n/k}(k-d, n-d)$ . In contrapositive, we will prove that if  $\delta_d(H) \geq f_0^{n/k}(k-d, n-d)$  then  $H$  has a fractional perfect matching, from which it follows, by definition, that  $f_d(k, n) \leq f_0^{n/k}(k-d, n-d)$ .

Let an  $n$ -vertex  $k$ -graph  $H$  satisfy  $\nu^*(H) < n/k$ , that is, have no fractional perfect matching. As  $\tau^*(H) = \nu^*(H)$ , there is a function  $w : V \rightarrow [0, 1]$  such that  $\sum_{v \in V} w(v) < n/k$  and, for every  $e \in H$ , we have  $\sum_{v \in e} w(v) \geq 1$ . We can replace  $H$  with the  $k$ -graph whose edge set consists of every  $k$ -tuple of vertices on which  $w$  totals to at least one.

Formally, for every weight function  $w : V \rightarrow [0, 1]$  define

$$H_w := \left\{ e \in \binom{V}{k} : \sum_{v \in e} w(v) \geq 1 \right\}.$$

For a given weight function  $w$ , suppose  $L$  is the set of  $d$  vertices with the smallest weights. Without loss of generality, we may assume that the  $d$  lowest values of  $w(x)$  are all equal to each other, since otherwise we could replace them by their average. (Obviously, this modification would not change  $\sum_{v \in V} w(v)$  nor the set  $L$ .) Note that the minimum  $d$ -degree  $\delta_d(H_w) = \min_{S \subset \binom{V}{d}} \deg_H(S)$  is achieved when  $S = L$ . Let  $H(L)$  be the neighborhood of  $L$  in  $H_w$ , that is a  $(k-d)$ -graph on the vertex set  $V \setminus L$  and with the edge set

$$\left\{ S \in \binom{V-L}{k-d} : S \cup L \in E(H_w) \right\}.$$

Then  $|H(L)| = \delta_d(H_w)$  and it remains to prove that  $\tau^*(H(L)) < n/k$ .

Let  $w_0 = \min_{v \in V} w(v)$  and observe that  $w_0 < 1/k$ . If  $w_0 > 0$ , apply the linear map

$$w' = \frac{w - w_0}{1 - kw_0}.$$

Then, still  $\sum_{v \in V} w'(v) < n/k$  and  $H_w = H_{w'}$ . Moreover, for every  $v \in L$ , we have  $w'(v) = 0$ . It follows that the function  $w'$  restricted to the set  $V \setminus L$  is a fractional vertex cover of  $H(L)$  and so  $\nu^*(H(L)) = \tau^*(H(L)) < n/k$ , which completes the proof of Proposition 1.1.  $\blacksquare$

**Proof of Theorem 1.2:** As explained earlier,  $f_0^{n/k}(k-d, n-d) = n/k$  holds trivially for  $d = k-1$  and together with Proposition 1.1 implies the theorem in this case. For  $d = k-2$ , we apply Proposition 1.1 together with the case  $l = 2$  of the fractional Erdős Conjecture 1.4 (as mentioned earlier, it follows asymptotically from [6]). For  $d = k-3$  and  $d = k-4$ , we use Proposition 1.1 and Corollary 1.3 proved in the previous section.  $\blacksquare$

**Remark 3.1** Consider a restricted version of Samuels' problem to minimize  $\mathbb{P}(X_1 + \dots + X_l < 1)$  under the *additional* assumption that all random variables are identically distributed. Our proofs indicate that under this regime, for a given  $l \geq 5$  and  $\mu_1 = \dots = \mu_l = x \leq \frac{1}{l+1}$ , if

$$\mathbb{P}(X_1 + \dots + X_l < 1) \geq (1 + o(1))(1 - x)^l$$

then Theorem 1.2 would hold for all  $k \geq l+1$  and  $d = k-l$ .

## 4 Constructing integer matchings from fractional ones

In this section, we will prove Theorem 1.1. An indispensable tool in our proof is the Strong Absorbing Lemma 4.1 from [10] (see Lemma 10 therein). This lemma provides a sufficient condition on degrees and co-degrees of a hypergraph ensuring the existence of a small and powerful matching which, by “absorbing” vertices, creates a perfect matching from any nearly perfect matching.

**Lemma 4.1** *For all  $\gamma > 0$  and integers  $k > d > 0$  there is an  $n_0$  such that for all  $n > n_0$  the following holds: suppose that  $H$  is a  $k$ -graph on  $n$  vertices with  $\delta_d(H) \geq (1/2 + 2\gamma) \binom{n-d}{k-d}$ , then there exists a matching  $M := M_{abs}$  in  $H$  such that*

(i)  $|M| < \gamma^k n/k$ , and

(ii) for every set  $W \subset V \setminus V(M)$  of size at most  $|W| \leq \gamma^{2k} n$  and divisible by  $k$  there exists a matching in  $H$  covering exactly the vertices of  $V(M) \cup W$ .

Equipped with this lemma we can practically reduce our task to finding an almost perfect matching in a suitable subhypergraph of  $H$ . Here is an outline of our proof of Theorem 1.1. Assume that there exists a constant  $0 < c^* < 1$  such that  $f_d(k, n) \sim c^* \binom{n-d}{k-d}$ . For any  $\alpha > 0$  consider a  $k$ -graph  $H$  on  $n$  vertices, where  $n$  is sufficiently large, with

$$\delta_d(H) \geq (c + \alpha) \binom{n-d}{k-d},$$

where  $c = \max\{\frac{1}{2}, c^*\}$ . Our goal is to show that  $H$  contains a perfect matching.

Set  $\gamma = \alpha/2$  and  $\varepsilon = \gamma^{2k}$ . The proof consists of three steps.

1. Find an absorbing matching  $M_{abs}$  satisfying properties (i) and (ii) of Lemma 4.1. Set  $H' = H \setminus V(M_{abs})$  and note that when  $n$  is sufficiently large,

$$\delta_d(H') \geq \delta_d(H) - \left( \binom{n-d}{k-d} - \binom{n-d-\varepsilon n}{k-d} \right) \geq (c + \alpha/2) \binom{n-d}{k-d} = (c + \gamma) \binom{n-d}{k-d}.$$

2. Find a matching  $M_{alm}$  in  $H'$  such that  $|V(M_{alm})| \geq (1 - \varepsilon)|V(H')|$ , and thus,  $|V(M_{alm} \cup M_{abs})| \geq (1 - \varepsilon)n$ .
3. Extend  $M_{alm} \cup M_{abs}$  to a perfect matching of  $H$  by using the absorbing property (ii) of  $M_{abs}$  with respect to  $W = V(H') \setminus V(M_{alm})$ .

Now come the details of the proof. The Strong Absorbing Lemma provides an absorbing matching  $M_{abs}$ , so Steps 1 and 3 are clear. Hence to complete the proof of Theorem 1.1 it remains to explain Step 2. One possible approach to find an almost perfect matching in  $H'$  is via the weak hypergraph regularity lemma. Our proof, however, is based on Theorem 1.1 in [8]. Recall that the codegree of a pair of vertices in a hypergraph is the number of edges containing this pair. An immediate corollary of that theorem asserts the existence of an almost perfect matching in any nearly regular  $k$ -graph in which all codegrees are much smaller than the vertex degrees. (See Remark after Theorem 1.1 in [8] or Chapter 4.7 of [2]). Here we formulate this corollary as the following lemma in which  $\Delta_2(H)$  denotes the maximum codegree in  $H$ .

**Lemma 4.2** *For every integer  $k \geq 2$  and a real  $\varepsilon > 0$ , there exists  $\tau = \tau(k, \varepsilon)$ ,  $d_0 = d_0(k, \varepsilon)$  such that for every  $n \geq D \geq d_0$  the following holds.*

*Every  $k$ -uniform hypergraph on a set  $V$  of  $n$  vertices which satisfies the following conditions:*

1.  $(1 - \tau)D < \deg_H(v) < (1 + \tau)D$  for all  $v \in V$ , and
2.  $\Delta_2(H) < \tau D$

*contains a matching  $M_{alm}$  covering all but at most  $\varepsilon n$  vertices.*

Hence, Step 2 above reduces to finding a spanning subhypergraph  $H''$  of  $H'$  satisfying the assumptions of Lemma 4.2 with  $\varepsilon = \gamma^{2k}$  and other parameters  $\tau, D, a$  to be suitably chosen. Indeed, the following claim is all we need to complete the proof of Theorem 1.1. For convenience, we set  $n := |V(H')|$ . Recall that  $c = \max\{\frac{1}{2}, c^*\}$  where  $c^*$  comes from the threshold which guarantees the existence of fractional perfect matchings.

**Claim 4.1** *For sufficiently large  $n$ , any  $k$ -graph  $H'$  on  $n$  vertices satisfying  $\delta_d(H') \geq (c + \gamma) \binom{n-d}{k-d}$  contains a spanning subhypergraph  $H''$ , such that for all  $v \in V(H'')$  we have  $\deg_{H''}(v) \sim n^{0.2}$  while  $\Delta_2(H'') \leq n^{0.1}$ .*

Consequently for every  $k \geq 2$ ,  $\varepsilon > 0$ , the subhypergraph  $H''$  satisfies the assumptions of Lemma 4.2 with  $D = n^{0.2}$ , and any  $\tau > 0$ . We obtained the following result as an immediate corollary, which asserts the validity of Step 2 and completes our proof of Theorem 1.1.

**Corollary 4.1**  *$H'$  contains an almost perfect matching covering at least  $(1 - \varepsilon)|V(H')|$  vertices.*

In the proof of Claim 4.1, the following well-known concentration results (see, for example [2], Appendix A, and Theorem 2.8, inequality (2.9) and (2.11) in [12]) will be used several times. We denote by  $Bi(n, p)$  a binomial random variable with parameters  $n$  and  $p$ .

**Lemma 4.3** *(Chernoff Inequality for small deviation) If  $X = \sum_{i=1}^n X_i$ , each random variable  $X_i$  has Bernoulli distribution with expectation  $p_i$ , and  $\alpha \leq 3/2$ , then*

$$\mathbb{P}(|X - \mathbb{E}X| \geq \alpha \mathbb{E}X) \leq 2e^{-\frac{\alpha^2}{3} \mathbb{E}X} \quad (6)$$

*In particular, when  $X \sim Bi(n, p)$  and  $\lambda < \frac{3}{2}np$ , then*

$$\mathbb{P}(|X - np| \geq \lambda) \leq e^{-\Omega(\lambda^2/(np))} \quad (7)$$

**Lemma 4.4** *(Chernoff Inequality for large deviation) If  $X = \sum_{i=1}^n X_i$ , each random variable  $X_i$  has Bernoulli distribution with expectation  $p_i$ , and  $x \geq 7 \mathbb{E}X$ , then*

$$\mathbb{P}(X \geq x) \leq e^{-x} \quad (8)$$

**Proof of Claim 4.1:** The desired subhypergraph  $H''$  is obtained via two rounds of randomization. As a preparation toward the first round,  $R$  is obtained by choosing every vertex randomly and independently with probability  $p = |V'|^{-0.9} = n^{-0.9}$ . Then  $|R|$  is a binomial random variable with expectation  $n^{0.1}$ . By inequality (7),  $|R| \sim n^{0.1}$  with probability  $1 - e^{-\Omega(n^{0.1})}$ .

Fix a subset  $D \subseteq V'$  of size  $d$  and let  $DEG_D$  be the number of edges  $f \in H'$  such that  $D \subset f$  and  $f \setminus D \subseteq R$ , which is the number of edges  $e$  in the link graph  $H[D]$  with all of its vertices in the random set  $R$ . Therefore  $DEG_D = \sum_{e \in H[D]} X_e$ , where  $X_e = 1$  if  $e$  is in  $R$  and 0 otherwise. We have

$$\begin{aligned} \mathbb{E}(DEG_D) &= \deg_{H'}(D) \times (n^{-0.9})^{k-d} \geq (c + \alpha/2) \binom{n-d}{k-d} n^{-0.9(k-d)} \\ &\geq (c + \alpha/3) \binom{|R| - d}{k-d} = \Omega(n^{0.1(k-d)}) \end{aligned}$$

For two distinct intersecting edges  $e_i, e_j$  with  $|e_i \cap e_j| = l$  for  $1 \leq l \leq k - d - 1$ , the probability that both of them are in  $R$  is

$$\mathbb{P}(X_{e_i} = X_{e_j} = 1) = p^{2(k-d)-l}$$

For fixed  $l$ , there are at most  $\binom{n-d}{k-d}$  choices for  $e_i$  in the link graph  $H[D]$ ,  $\binom{k-d}{l}$  ways to choose the intersection  $L = e_i \cap e_j$  of size  $l$ , and  $\binom{(n-d)-(k-d)}{k-d-l}$  options for  $e_j \setminus L$ . Therefore,

$$\begin{aligned} \Delta &= \sum_{e_i \cap e_j \neq \emptyset} \mathbb{P}(X_{e_i} = X_{e_j} = 1) \leq \sum_{l=1}^{k-d-1} p^{2(k-d)-l} \binom{n-d}{k-d} \binom{k-d}{l} \binom{n-k}{k-d-l} \\ &\leq \sum_{l=1}^{k-d-1} p^{2(k-d)-l} O(n^{2(k-d)-l}) = O(n^{0.1(2(k-d)-1)}) \end{aligned}$$

By Janson's inequality (see Theorem 8.7.2 in [2]),

$$\mathbb{P}(DEG_D \leq (1 - \alpha/12)\mathbb{E}(DEG_D)) \leq e^{-\Omega((\mathbb{E}X)^2/\Delta)} \sim e^{-\Omega(n^{0.1})}$$

Therefore by the union bound, with probability  $1 - n^d e^{-\Omega(n^{0.1})}$ , for all subsets  $D \subseteq V'$  of size  $d$ , we have

$$DEG_D > (1 - \alpha/12)\mathbb{E}(DEG_D) \geq (c + \alpha/4) \binom{|R| - d}{k-d}.$$

Take  $n^{1.1}$  independent copies of  $R$  and denote them by  $R^i$ ,  $1 \leq i \leq n^{1.1}$ , and the corresponding random variables by  $DEG_D^{(i)}$ , where  $D \subseteq V'$ ,  $|D| = d$ , and  $i = 1, \dots, n^{1.1}$ . Since  $|R_i| \sim n^{0.1}$  with probability  $1 - e^{-\Omega(n^{0.1})}$  for each  $i$ , the union bound ensures that  $|R_i| \sim n^{0.1}$  for every  $i = 1, \dots, n^{1.1}$  with probability  $1 - o(1)$ . Now for a subset of vertices  $S \subseteq V'$ , define the random variable

$$Y_S = |\{i : S \subseteq R^i\}|.$$

Note that the random variables  $Y_S$  have binomial distributions  $Bi(n^{1.1}, n^{-0.9|S|})$  with expectations  $n^{1.1-0.9|S|}$ . In particular, for every vertex  $v \in V'$ ,  $Y_{\{v\}} \sim Bi(n^{1.1}, n^{-0.9})$  and  $\mathbb{E}Y_{\{v\}} = n^{0.2}$ . Hence, by inequality (7), taking  $\lambda = n^{0.15}$ ,

$$\mathbb{P}(|Y_{\{v\}} - n^{0.2}| > n^{0.15}) \leq e^{-\Omega((n^{0.15})^2/n^{0.2})} = e^{-\Omega(n^{0.1})}$$

Therefore a.a.s  $|Y_{\{v\}} - n^{0.2}| \leq n^{0.15}$  for every vertex  $v \in V'$ .

Further, let

$$Z_2 = \left| \left\{ \{u, v\} \in \binom{V'}{2} : Y_{\{u, v\}} \geq 3 \right\} \right|.$$

Then

$$\mathbb{E}Z_2 < n^2(n^{1.1})^3(n^{-0.9})^6 = n^{-0.1}.$$

Therefore by Markov's inequality,

$$\mathbb{P}(Z_2 = 0) = 1 - \mathbb{P}(Z_2 \geq 1) \geq 1 - \mathbb{E}Z_2 > 1 - n^{-0.1}$$

This implies that a.a.s every pair of vertices  $\{u, v\}$  is contained in at most two subhypergraphs  $R^i$ .

Finally, for  $k \geq 3$ , let

$$Z_k = \left| \left\{ S \in \binom{V'}{k} : Y_S \geq 2 \right\} \right|.$$

Then,

$$\mathbb{E}Z_k < n^k (n^{1.1})^2 (n^{-0.9})^{2k} = n^{k+2.2-1.8k} \leq n^{-0.2}$$

Similarly,

$$\mathbb{P}(Z_k = 0) > 1 - n^{-0.2}$$

The latter implies that a.a.s. the induced subhypergraphs  $H[R^i]$ ,  $i = 1, \dots, n^{1.1}$ , are pairwise edge-disjoint. Summarizing, we can choose the sets  $R^i$ ,  $1 \leq i \leq n^{1.1}$  in such a way that

- (i) for every  $v \in V'$ ,  $Y_{\{v\}} \sim n^{0.2}$ ,
- (ii) every pair  $\{u, v\} \subset V'$  is contained in at most *two* sets  $R^i$ ,
- (iii) every edge  $e \in H$  is contained in at most *one* set  $R^i$ ,
- (iv) for all  $i = 1, \dots, n^{1.1}$ , we have  $|R^i| \sim n^{0.1}$ , and
- (v) for all  $i = 1, \dots, n^{1.1}$  and all  $D \subseteq V'$ ,  $|D| = d$ , we have  $DEG_D^{(i)} \geq (c + \alpha/4) \binom{|R^i| - d}{k - d}$ .

Let us fix a sequence  $R^i$ ,  $1 \leq i \leq n^{1.1}$ , satisfying (i)-(v) above.

Our assumption that  $f_d(k, n) \sim c^* \binom{n-d}{k-d}$  holds for all sufficiently large values of  $n$ , in particular with  $n$  replaced by  $|R^i| \sim n^{0.1}$ . Thus, we have

$$f_d(k, |R^i|) \sim c^* \binom{|R^i| - d}{k - d},$$

and, by condition (v) above, we conclude that

$$\delta_d(H[R^i]) \geq (c + \alpha/4) \binom{|R^i| - d}{k - d} > f_d(k, |R^i|).$$

Consequently, by the definition of  $f_d$ , there exists a fractional perfect matchings  $w^i$  in every subhypergraph  $H[R^i]$ ,  $i = 1, \dots, n^{1.1}$ .

Now comes the second round of randomization. Let  $H^* = \bigcup_i H[R^i]$ . We select a generalized binomial subhypergraph  $H''$  of  $H^*$  by independently choosing each edge  $e$  with probability  $w^{i_e}(e)$ , where  $i_e$  is the index  $i$  such that  $e \in H[R^i]$ . Recall that property (iii) ensures that every edge is contained in at most one hypergraph  $R^i$ , which guarantees the uniqueness of  $i_e$ . We are going to verify our claim by showing  $\deg_{H''}(v) \sim n^{0.2}$  for any vertex  $v$ , while  $\Delta_2(H'') \leq n^{0.1}$ .

Let  $I_v = \{i : v \in R^i\}$  and recall that  $|I_v| = Y_{\{v\}} \sim n^{0.2}$  by (i). For every  $v \in V'$  the set  $E_v$  of edges  $e \in H^*$  containing  $v$  can be partitioned into  $|I_v|$  parts  $E_v^i = \{e \in E_v \cap H[R^i]\}$ . Recall that  $w^i$  is a perfect matching, and thus  $\sum_{e \in E_v^i} w^i(e) = 1$ . For every  $v \in V'$  the random variable  $D_v = \deg_{H''}(v)$  is

equal to  $\sum_{i \in I_v} \sum_{e \in E_v^i} X_e$ , where  $X_e$  are independent random variables having Bernoulli distribution with expectation  $w^{ie}(e)$ . Therefore  $D_v$  is generalized binomial with expectation

$$\mathbb{E}D_v = \sum_{e \in E_v} w^{ie}(e) = \sum_{i \in I_v} \left( \sum_{e \in E_v^i} w^i(e) \right) = \sum_{i \in I_v} 1 \sim n^{0.2}.$$

Hence by Chernoff's inequality (6),

$$\mathbb{P}(|D_v - n^{0.2}| \geq \alpha n^{0.2}) \leq 2e^{-\frac{\alpha^2}{3} n^{0.2}}$$

Set  $\alpha = n^{-0.05}$ , then  $|D_v - n^{0.2}| \leq n^{0.15}$  with probability  $1 - O(e^{-n^{0.1}})$ . Taking a union bound over all the  $n$  vertices, we conclude that a.a.s. for all  $v \in V'$  we have  $D_v \sim n^{0.2}$ .

Moreover, for all pairs  $u, v \in V'$  the random variable  $D_{u,v} = \text{deg}_{H''}(u, v)$  is also generalized binomial with expectation

$$\mathbb{E}D_{u,v} = \sum_{e \in E_u \cap E_v} w^{ie}(e) = \sum_{i \in I_u \cap I_v} \left( \sum_{e \in E_u^i \cap E_v^i} w^i(e) \right) \leq |I_u \cap I_v| \leq 2$$

by (ii). Hence, again by Chernoff's inequality (8) for large deviations, when  $n$  is sufficiently large,

$$\mathbb{P}(D_{u,v} \geq n^{0.1}) \leq e^{-n^{0.1}}$$

Once again taking the union bound ensures that a.a.s. for every pair of vertices  $u, v \in V'$ ,  $D_{u,v} \leq n^{0.1}$ . ■

## 5 An applications in distributed storage allocation

The following model of distributed storage has been studied in information theory [17, 20, 28]. A file is split into multiple chunks, replicated redundantly and stored in a distributed storage system with  $n$  nodes. Suppose the amount of data to be stored in each node  $i$  is equal to  $x_i$ , where the size of the whole file is normalized to 1. In reality, because there is limited storage space or transmission bandwidth, we require that the total amount of data stored does not exceed a given budget  $T$ , i.e.  $x_1 + \dots + x_n \leq T$ . At the time of retrieval, we attempt to recover the whole file by accessing only the data stored in a subset  $R$  of  $r$  nodes which is chosen uniformly at random. It is known that there always exists a coding scheme such that we can recover the file whenever the total amount of data accessed is at least 1. Our goal is to find an optimal allocation  $(x_1, \dots, x_n)$  in order to maximize the probability of successful recovery. This problem can be reformulated as follows.

**Question 5.1** For a nonnegative sequence  $(x_1, \dots, x_n)$ , let

$$\Phi(x_1, \dots, x_n) = \left| \left\{ S \subseteq [n], |S| = r \text{ such that } \sum_{i \in S} x_i \geq 1 \right\} \right|.$$

Then the probability of successful recovery of the file equals

$$\frac{\Phi(x_1, \dots, x_n)}{\binom{n}{r}}.$$

Given integers  $n \geq r \geq 1$  and a real number  $T > 0$ , determine

$$F^T(r, n) = \max_{\sum x_i = T, x_i \geq 0 \forall i} \Phi(x_1, \dots, x_n).$$

and find an allocation optimizing  $F^T(r, n)$ .

If the total budget  $T$  is at least  $n/r$  then, by setting all  $x_i = T/n \geq 1/r$  for all  $i$ , we can recover the original file from any subset of size  $r$ . So,  $F^T(r, n) = \binom{n}{r}$  for  $T \geq n/r$ . For  $T < n/r$ , let  $w(i) = x_i$  be a weight function from  $V = [n]$  to  $\mathbb{R}$ . Then by the definition of the threshold  $r$ -uniform hypergraph  $H_w^1$  from Section 3, the edges of  $H_w^1$  correspond to the  $r$ -subsets  $S$  such that  $\sum_{i \in S} x_i \geq 1$ . Thus, it is easy to see that the fractional matching number of  $H_w^1$  satisfies

$$\nu^*(H_w^1) = \tau^*(H_w^1) \leq \sum_{i=1}^n w(i) = \sum_{i=1}^n x_i \leq T$$

while

$$\Phi(x_1, \dots, x_n) = |H_w^1|.$$

Therefore,  $F^T(r, n)$  is the maximum number of edges in an  $r$ -uniform hypergraph on  $n$  vertices with fractional matching number at most  $T$ . As such  $F^T(r, n)$  differs from  $f_0^T(r, n)$  only in that the latter has the strict inequality  $\nu^*(H) < T$  in its definition. But, of course, we have  $f_0^T(r, n) \leq F^T(r, n) \leq f_0^{T+1}(r, n)$ , and so  $F^T(r, n) \sim f_0^T(r, n)$  as  $n \rightarrow \infty$ .

Hence, Question 5.1 is asymptotically equivalent to the fractional Erdős Conjecture 1.4. As mentioned in the introduction, it follows from the Erdős-Gallai theorem [6] that

$$F^T(2, n) \sim f_0^T(2, n) \sim m_0^T(2, n) \sim \max \left\{ \binom{2T}{2}, \binom{n}{2} - \binom{n-T}{2} \right\}.$$

An easy calculation shows that the above maximum equals the first term if  $\frac{2}{5}n \leq T \leq \frac{1}{2}n$ , and the corresponding optimal graph is a clique of size  $2T$ . This means that, asymptotically, an optimal allocation is  $x_1 = \dots = x_{2T} = 1/2$  and  $x_{2T+1} = \dots = x_n = 0$ . On the other hand, if  $T < \frac{2}{5}n$ , an optimal allocation is  $x_1 = \dots = x_T = 1$  and  $x_{T+1} = \dots = x_n = 0$ .

For general  $r \geq 3$ , if Conjecture 1.4 is true, then

$$F^T(r, n) \sim \max \left\{ \binom{rT}{r}, \binom{n}{r} - \binom{n-T}{r} \right\}.$$

The bounds are achieved when  $H$  is a clique or a complement of clique. A corresponding (asymptotically) optimal storage allocation is  $x_1 = \dots = x_{rT} = 1/r, x_{rT+1} = \dots = x_n = 0$  or  $x_1 = \dots = x_T = 1, x_{T+1} = \dots = x_n = 0$ , respectively. Corollary 2.1 and Remark 2.1 assert that



for  $r = 3$  and  $T < 0.277 n$ , as well as for  $r = 4$  and  $T < 0.217 n$ , the latter is an optimal allocation. Moreover, if Samuels' conjecture 2.1 holds for all the remaining  $r \geq 5$ , then  $x_1 = \dots = x_T = 1, x_{T+1} = \dots = x_n = 0$  is always an asymptotic optimal allocation whenever  $T < n/(r+1)$ . Erdős [5] proved Conjecture 1.3 for all  $T < n/(2r^3)$ . Recently, the authors of [11] extended the range for which this conjecture holds to  $T = O(n/r^2)$ . Therefore, in this range,  $F^T(r, n)$  is achieved by the complement of a clique and an optimal allocation is also known to be  $x_1 = \dots = x_T = 1, x_{T+1} = \dots = x_n = 0$ .

## 6 Concluding Remarks

- We have studied sufficient conditions on the minimum  $d$ -degree which guarantee that a uniform hypergraph has a perfect matching or perfect fractional matching. We proved that if  $f_d(k, n) \sim c^* \binom{n}{k}$ , then  $m_d(k, n) \sim \max\{c^*, 1/2\} \binom{n}{k}$ . Therefore in order to determine the asymptotic behavior of the minimum  $d$ -degree ensuring existence of a perfect matching, we can instead study the presumably easier question for fractional matchings. Using this approach we showed, in particular, that  $m_1(5, n) \sim \left(1 - \frac{4^4}{5^4}\right) \binom{n-1}{4}$ .
- An intriguing problem which remains open is the conjecture by Erdős which states that the maximum number of edges in a  $k$ -uniform hypergraph  $H$  on  $n$  vertices with matching number smaller than  $s$  is exactly

$$\max \left\{ \binom{ks-1}{k}, \binom{n}{k} - \binom{n-s+1}{k} \right\}.$$

The fractional version of Erdős conjecture is also very interesting. In its asymptotic form it says that if  $H$  is an  $l$ -uniform  $m$ -vertex hypergraph with fractional matching number  $\nu^*(H) = xm$ , where  $0 \leq x < 1/l$ , then

$$|H| \leq (1 + o(1)) \max \{(lx)^l, 1 - (1-x)^l\} \binom{m}{l}.$$

In Section 2 we showed that the fractional Erdős conjecture is related to a probabilistic conjecture of Samuels. This conjecture, if proved, will provide a solution to the fractional version of Erdős problem for the range  $x \leq \frac{1}{l+1}$ . It will also lead to the asymptotics of  $m_d(k, n)$  and  $f_d(k, n)$  for arbitrary  $k \geq d+1$  and  $d \geq 1$ .

- As it turns out, matchings and fractional matchings also have some interesting applications in information theory. In particular, the uniform model of distributed storage allocation considered in [28] leads to a question which is asymptotically equivalent to the fractional version of Erdős' problem. In [17], the set of accessed nodes,  $R$ , is given by taking each node randomly and independently with probability  $p$ . It would be interesting to see if our techniques can be applied to study this binomial model too.

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